

第五章、极限定理

§5.1 Bernoulli 试验场合的极限定理

- 假设 $X = X_1, X_2, \dots$ 独立同分布(i.i.d.),

$$P(X = 1) = p, \quad P(X = 0) = q = 1 - p.$$

$$S_n = X_1 + \cdots + X_n.$$

- $P(S_n = k) = C_n^k p^k q^{n-k}, k = 0, 1, \dots, n.$

- 频率 $\xi_n = \frac{1}{n} S_n = \bar{X}$:

$$E S_n = np, \quad \text{var}(S_n) = npq.$$

$$E \xi_n = p, \quad \text{var}(\xi_n) = \frac{n \cdot \text{var}(X)}{n^2} = \frac{1}{n} \text{var}(X) = \frac{1}{n} pq.$$

- (1713年) Bernoulli 大数定律(5.1.14):

$$P\left(\left|\underbrace{\frac{S_n}{n}}_{\text{在 } n \rightarrow \infty \text{ 时}} - \underbrace{p}_{\text{固定}}\right| \geq \varepsilon\right) \leq \frac{1}{\varepsilon^2} \text{var}\left(\frac{S_n}{n}\right) = \frac{1}{\varepsilon^2} \cdot \frac{\text{var}(X)}{n} \rightarrow 0.$$

- 如果 $\forall \varepsilon > 0$ 都有

$$\lim_{n \rightarrow \infty} P(|\xi_n - \xi| \geq \varepsilon) = 0.$$

那么, 称 ξ_n 依概率收敛到 ξ , 记为 $\xi_n \xrightarrow{P} \xi$. (定义5.2.3)

- $S_n^* = \frac{S_n - np}{\sqrt{npq}}$: $x = \varepsilon / \sqrt{pq}$,

$$P\left(\left|\frac{S_n - np}{\sqrt{n}}\right| \geq \varepsilon\right) = P(|S_n^*| \geq x) \leq \frac{1}{x^2}, \quad \forall x > 0.$$

- (1733, 1778) De Moivre-Laplace 中心极限定理(定理5.1.1):

$$P(a < S_n^* \leq b) \rightarrow \int_a^b \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx, \quad \forall a < b.$$

- $x_k = \frac{k-np}{\sqrt{npq}}$, $\mathbb{K} = [np + a\sqrt{npq}, np + b\sqrt{npq}]$, $\Delta x = \frac{1}{\sqrt{npq}}$.

$$\frac{P(S_n = k)}{p_Z(x_k)\Delta x} = \frac{P(S_n = k)}{\frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}x_k^2} \frac{1}{\sqrt{npq}}} \xrightarrow{\mathbb{K}} 1.$$

- 进一步,

$$\lim_n \frac{\sum_{k \in \mathbb{K}} P(S_n = k)}{\int_a^b \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx} = \lim_n \frac{\sum_{k \in \mathbb{K}} P(S_n = k)}{\sum_{k \in \mathbb{K}} p_Z(x_k)\Delta x} = 1.$$

- a 可为 $-\infty$; b 可为 ∞ .

$$P(\textcolor{red}{S_n^*} \leq x) \rightarrow \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} dz = \Phi(x), \quad \forall x \in \mathbb{R}.$$

- 假设 $\xi \sim N(0, 1)$. 如果 $\forall x$ 都有

$$\lim_{n \rightarrow \infty} P(\textcolor{red}{\xi_n} \leq x) = P(\textcolor{blue}{\xi} \leq x).$$

那么, 称 ξ_n 依分布收敛于 ξ , 记为 $\xi_n \xrightarrow{d} \xi$.

§5.2 之依概率收敛, §5.1 & §5.3 之弱大数定律

- 如果 $\forall \varepsilon > 0$ 都有

$$\lim_{n \rightarrow \infty} P(|\xi_n - \xi| \geq \varepsilon) = 0.$$

那么, 称 ξ_n 依概率收敛到 ξ , 记为 $\xi_n \xrightarrow{P} \xi$. (定义5.2.3)

- 如果

$$\lim_{n \rightarrow \infty} E|\xi_n - \xi|^r = 0.$$

那么, 称 ξ_n r 阶(平均)收敛到 ξ , 记为 $\xi_n \xrightarrow{r} \xi$. (定义5.2.4)

- $\xi_n \xrightarrow{r} \xi \Rightarrow \xi_n \xrightarrow{P} \xi$. (定理5.2.8)

$$P(|\xi_n - \xi| \geq \varepsilon) \leq \frac{1}{\varepsilon^r} E|\xi_n - \xi|^r.$$

- “ $\xi_n \xrightarrow{P} \xi \Rightarrow \xi_n \xrightarrow{r} \xi$ ” 不成立. (例5.2.4)

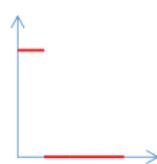
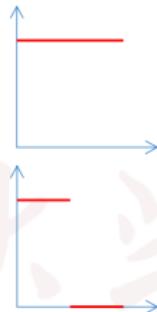
$\Omega = (0, 1]$, 几何概率.

$$\xi_n = n^{\frac{1}{r}} \times 1_{(0, \frac{1}{n}]},$$

$$\xi \equiv 0.$$

- $\xi_n \xrightarrow{P} \xi$ & $E|\xi_n|^r \rightarrow E|\xi|^r \Rightarrow \xi_n \xrightarrow{r} \xi$.

(证明不要求)



- 弱大数定律(Weak Law of Large Numbers, WLLN)形式:
若 X_1, X_2, \dots 满足 ***, 则

$$\frac{S_n}{b_n} - a_n \xrightarrow{P} 0.$$

- WLLN 的证明方法:

$$P\left(\left|\frac{S_n}{b_n} - a_n\right| \geq \varepsilon\right) \leq \frac{1}{\varepsilon^r} E\left|\frac{S_n}{b_n} - a_n\right|^r, \quad r > 0.$$

- Chebyshev's WLLN: 两两不相关, $\text{var}(X_i) \leq M, \forall i$, 则

$$\frac{S_n - ES_n}{n} \xrightarrow{P} 0. \quad (5.1.12)$$

- Markov's WLLN: $\text{var}(S_n) = o(n^2)$, 则

$$\frac{S_n - ES_n}{n} \xrightarrow{P} 0. \quad (5.1.12)$$

- 又若 $\frac{ES_n}{n} \rightarrow a$, 则 $\frac{S_n}{n} \xrightarrow{P} a$.
- 习题五、26(4).

$$\xi_n \xrightarrow{P} \xi, \quad \eta_n \xrightarrow{P} \eta \quad \Rightarrow \quad \xi_n + \eta_n \xrightarrow{P} \xi + \eta.$$

证明: $\forall \varepsilon > 0$,

$$\begin{aligned} & P\left(\left|(\xi_n + \eta_n) - (\xi + \eta)\right| \geq \varepsilon\right) \\ & \leq P\left(|\xi_n - \xi| \geq \frac{\varepsilon}{2}\right) + P\left(|\eta_n - \eta| \geq \frac{\varepsilon}{2}\right). \end{aligned}$$

- Bernoulli's WLLN (5.1.14), Poisson's WLLN (5.1.16).

习题五、44. $f : [0, 1] \rightarrow \mathbb{R}$ 连续. 则存在多项式 f_n 使得 $f_n \xrightarrow{[0,1]} f$.

- 设 X_1, X_2, \dots i.i.d., $P(X_1 = 1) = 1 - P(X_1 = 0) = x$. 则

$$\text{LLN: } \frac{S_n}{n} \rightarrow x \Rightarrow \textcolor{blue}{f}\left(\frac{S_n}{n}\right) \approx f(x) \text{(直观).}$$

- 取 f_n :

$$f_n(x) = E \textcolor{blue}{f}\left(\frac{S_n}{n}\right) = \sum_{k=0}^n f\left(\frac{k}{n}\right) C_n^k x^k (1-x)^{n-k}.$$

- $\forall \varepsilon > 0$, $\exists \delta$ 使得

$$\text{若 } |y - z| < \delta, \quad \text{则 } |f(y) - f(z)| \leq \frac{1}{2}\varepsilon.$$

- 令 $A_n = \left\{ \left| \frac{S_n}{n} - x \right| \geq \delta \right\} = \{|S_n - nx| \geq n\delta\}$. 则

$$f_n(x) - f(x) = E \left(\textcolor{blue}{f} \left(\frac{S_n}{n} \right) - f(x) \right) (1_{A_n} + 1_{A_n^c}).$$

- 取 $C = 2 \max_{x \in [0,1]} |f(x)|$, 则

$$|f_n(x) - f(x)| \leq C P(A_n) + \frac{1}{2}\varepsilon.$$

- $\sigma^2 = x(1-x) \leq \frac{1}{4}$: 当 $n \geq N$ 时,

$$\begin{aligned} |f_n(x) - f(x)| &\leq C \frac{n\sigma^2}{n^2\delta^2} + \frac{1}{2}\varepsilon \\ &\leq \frac{C}{4n\delta^2} + \frac{1}{2}\varepsilon \leq \varepsilon, \quad \forall x \in [0, 1]. \end{aligned}$$

有界收敛定理: 假设 $\xi_n \xrightarrow{P} \xi$, 且 $P(|\xi_n| \leq M) = 1, \forall n$. 那么,

$$\lim_{n \rightarrow \infty} E\xi_n = E\xi.$$

- 记 $A_n = \{|\xi_n - \xi| > \varepsilon\}$.

$$\begin{aligned} E|\xi_n - \xi| &= E|\xi_n - \xi| \cdot 1_{A_n} + E|\xi_n - \xi| \cdot 1_{A_n^c} \\ &\leq 2M \cdot P(A_n) + \varepsilon. \end{aligned}$$

- $\xi_n \xrightarrow{1} \xi$:

$$\limsup_{n \rightarrow \infty} E|\xi_n - \xi| \leq \varepsilon, \quad \forall \varepsilon > 0.$$

- $|E\xi_n - E\xi| \leq E|\xi_n - \xi| \rightarrow 0$.

定理5.3.1. 设 $X = X_1, X_2, \dots$ i.i.d., $E|X| < \infty$, 则 $\frac{S_n}{n} \xrightarrow{P} EX$.

- $E|X| < \infty \Rightarrow \lim_{x \rightarrow \infty} xP(|X| > x) = 0$.

$$\frac{x}{2} P(|X| > x) \leq \int_{x/2}^x P(|X| > y) dy \xrightarrow{x \rightarrow \infty} 0.$$

- $\frac{S_n}{n} - EX \cdot 1_{\{|X| \leq n\}} \xrightarrow{P} 0$. (见后面例*)
- $\mu_n = EX \cdot 1_{\{|X| \leq n\}} \rightarrow EX$:

$$\begin{aligned} E|X| \cdot 1_{\{|X| > n\}} &= \int_0^\infty P(|X| \cdot 1_{\{|X| > n\}} > y) dy \\ &= \int_0^\infty P(|X| > y \vee n) dy \\ &= \int_0^n P(|X| > n) dy + \int_n^\infty P(|X| > y) dy \rightarrow 0. \end{aligned}$$

例*. 假设 X_1, X_2, \dots i.i.d., $\lim_{x \rightarrow \infty} xP(|X| > x) = 0$, 则

$$\frac{S_n}{n} - EX \cdot 1_{\{|X| \leq n\}} \xrightarrow{P} 0.$$

● 截断:

$$T_n = X_1 \cdot 1_{\{|X_1| \leq n\}} + \cdots + X_n \cdot 1_{\{|X_n| \leq n\}}.$$

则, $P(S_n \neq T_n) \leq nP(X \neq X \cdot 1_{\{|X| \leq n\}}) = nP(|X| > n) \rightarrow 0$.

● 故,

$$P\left(\left|\frac{S_n}{n} - \mu_n\right| \geq \varepsilon\right) \leq P(S_n \neq T_n) + P\left(\left|\frac{T_n}{n} - \mu_n\right| \geq \varepsilon\right).$$

- $\text{var}(X \cdot 1_{\{|X| \leq n\}}) \leq EX^2 \cdot 1_{\{|X| \leq n\}} = o(n)$:

$$\begin{aligned} EX^2 \cdot 1_{\{|X| \leq n\}} &= \int_0^\infty P(\textcolor{red}{X^2} 1_{\{|X| \leq n\}} > \textcolor{red}{x}) dx \\ &= \int_0^{n^2} P(\sqrt{x} < |X| \leq n) dx \\ &\leq \int_0^n \underbrace{P(\textcolor{red}{y} < |X|)}_{\sim} \cdot 2y dy. \end{aligned}$$

- $P\left(\left|\frac{T_n}{n} - \mu_n\right| \geq \varepsilon\right) \rightarrow 0$:

$$P\left(\left|\frac{T_n}{n} - \mu_n\right| \geq \varepsilon\right) \leq \frac{1}{\varepsilon^2} \frac{\text{var}(T_n)}{n^2} = \frac{1}{\varepsilon^2} \frac{\text{var}(X \cdot 1_{\{|X| \leq n\}})}{n} \rightarrow 0.$$

例*(习题一38, 习题四6). 假设有 n 种券, 集齐时间为 S_n , 则

$$\frac{S_n}{n \ln n} \xrightarrow{P} 1.$$

- $S_n = X_1 + \cdots + X_n$:

$$X_k \sim G(p_k), \quad p_k = \frac{n - (k - 1)}{n}.$$

- 若 $X \sim G(p)$, 则

$$EX = \frac{1}{p}, \quad \text{var}(X) = \frac{1-p}{p^2} \leq \frac{1}{p^2}.$$

- 于是,

$$ES_n = n \sum_{\ell=1}^n \frac{1}{\ell} \approx n \ln n, \quad \text{var}(S_n) \leq n^2 \sum_{\ell=1}^n \frac{1}{\ell^2} \leq Cn^2.$$

- $\frac{ES_n}{n \ln n} \rightarrow 1$, $\text{var}(S_n) \leq Cn^2$:

$$P\left(\left|\frac{S_n - ES_n}{n \ln n}\right| \geq \varepsilon\right) \leq \frac{\text{var}(S_n)}{\varepsilon^2 n^2 (\ln n)^2} \rightarrow 0.$$

- $\frac{S_n - ES_n}{n \ln n} \xrightarrow{P} 0$, 故

$$\frac{S_n}{n \ln n} \xrightarrow{P} 1.$$

- 集齐一半的时间: $T_n = X_1 + \dots + X_{n/2}$.

$$ET_n \approx n \sum_{\ell=n/2}^n \frac{1}{\ell} \approx n \ln 2, \quad \text{var}(T_n) \leq n^2 \sum_{\ell=n/2}^n \frac{1}{\ell^2} \leq \delta_n n^2.$$

故, $P\left(\left|\frac{T_n - ET_n}{n \ln 2}\right| > \varepsilon\right) \leq \frac{\delta_n n^2}{\varepsilon n^2 (\ln 2)^2} \rightarrow 0, \quad \frac{T_n}{n \ln 2} \xrightarrow{P} 1.$

几乎必然收敛.

§5.2 之几乎必然收敛、§5.4 强大数定律

- 如果

$$P\left(\lim_{\substack{n \rightarrow \infty}} \xi_n = \xi\right) = 1.$$

那么, 称 ξ_n 几乎必然收敛到 ξ , 记为 $\xi_n \xrightarrow{\text{a.s.}} \xi$. (定义5.2.5)

- $\xi_n = \{\omega : \lim_{n \rightarrow \infty} \xi_n(\omega) = \xi(\omega)\}$.
- $\lim_{n \rightarrow \infty} \xi_n(\omega) = \xi(\omega)$ 不成立:

$\exists \varepsilon > 0, \forall N \geq 1, \exists n \geq N$ 使得 $|\xi_n(\omega) - \xi(\omega)| > \varepsilon$.

$\exists k \geq 1, \forall N \geq 1, \exists n \geq N$ 使得 $|\xi_n(\omega) - \xi(\omega)| > \frac{1}{k}$.

- 令 $A_{n,\varepsilon} = \{|\xi_n - \xi| > \varepsilon\}$, 则

$$\begin{aligned}\left\{\lim_{n \rightarrow \infty} \xi_n = \xi\right\}^c &= \bigcup_{\varepsilon > 0} \bigcap_{N \geq 1} \bigcup_{n \geq N} A_{n,\varepsilon} \\ &= \bigcup_{k \geq 1} \bigcap_{N \geq 1} \bigcup_{n \geq N} A_{n,\frac{1}{k}}.\end{aligned}$$

- 当 $\varepsilon \searrow 0$ 时,

$$A_\varepsilon := \bigcap_{N \geq 1} \bigcup_{n \geq N} A_{n,\varepsilon} \nearrow,$$

$$\star\star = \bigcup_{\varepsilon > 0} A_\varepsilon = \bigcup_{k \geq 1} A_{\frac{1}{k}} = \lim_{k \rightarrow \infty} A_{\frac{1}{k}}.$$

- 结论:

$$\xi_n \xrightarrow{\text{a.s.}} \xi \text{ iff } P(A_\varepsilon) = 0, \forall \varepsilon > 0 \text{ iff } P\left(A_{\frac{1}{k}}\right) = 0, \forall k \geq 1.$$

- 对任意事件列 A_1, A_2, \dots , 令

$$\{A_n \text{ i.o.}\} = \limsup_{n \rightarrow \infty} A_n := \bigcap_{N \geq 1} \bigcup_{n \geq N} A_n.$$

- Borel-Cantelli引理(引理5.4.1): 令 $s = \sum_{n=1}^{\infty} P(A_n)$.

- (1) 若 $s < \infty$, 则 $P(A_n \text{ i.o.}) = 0$.
- (2) 若 $s = \infty$, 且 A_1, A_2, \dots 相互独立, 则 $P(A_n \text{ i.o.}) = 1$.

- $\bigcup_{n \geq N} A_n \searrow: (1 - p \leq e^{-p},)$

$$P(A_n \text{ i.o.}) = \lim_{N \rightarrow \infty} P\left(\bigcup_{n \geq N} A_n\right).$$

$$\prod_{n \geq N} (1 - P(A_n)) \leq e^{-\sum_{n \geq N} P(A_n)} = 0.$$

总结:

- $\xi_n \xrightarrow{\text{a.s.}} \xi$ 的定义及等价条件:

$$P\left(\left\{\omega : \lim_{n \rightarrow \infty} \xi_n(\omega) = \xi(\omega)\right\}\right) = 1. \quad (\text{定义、应用、验证})$$

$$P(A_{n,\varepsilon} \text{ i.o.}) = 0, \quad \forall \varepsilon > 0; \quad (\text{验证})$$

$$P\left(A_{n,\frac{1}{k}} \text{ i.o.}\right) = 0, \quad \forall k \geq 1;$$

$$P\left(\bigcup_{n \geq N} A_{n,\varepsilon}\right) \searrow 0(N \rightarrow \infty), \quad \forall \varepsilon > 0.$$

其中, $A_{n,\varepsilon} = \{|\xi_n - \xi| > \varepsilon\}$.

- $\xi_n \xrightarrow{\text{a.s.}} \xi$ 的充分条件:

$$\sum_{n=1}^{\infty} P(A_{n,\varepsilon}) < \infty, \quad \forall \varepsilon > 0.$$

几乎必然收敛与其它收敛.

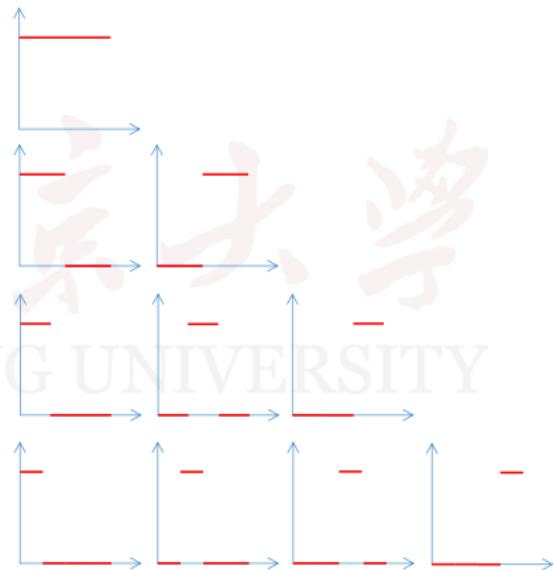
- $\xi_n \xrightarrow{\text{a.s.}} \xi \Rightarrow \xi_n \xrightarrow{P} \xi$. (定理5.4.1)

$$P\left(\bigcup_{n \geq N} A_{n,\varepsilon}\right) \geq P(A_{N,\varepsilon}),$$

$$A_{n,\varepsilon} = \{|\xi_n - \xi| > \varepsilon\}.$$

- “ $\xi_n \xrightarrow{P \text{或 } r} \xi \Rightarrow \xi_n \xrightarrow{\text{a.s.}} \xi$ ” 不成立.
(例5.4.1)

- “ $\xi_n \xrightarrow{\text{a.s.}} \xi \Rightarrow \xi_n \xrightarrow{r} \xi$ ” 不成立.
(例5.2.4)



- 若 $\xi_n \xrightarrow{P} \xi$, 则存在子列 $\{n_k\}$ 使得 $\xi_{n_k} \xrightarrow{\text{a.s.}} \xi$.

(1) $\exists n_1 \geq 1$ 使得

$$P(|\xi_n - \xi| > 1) < \frac{1}{2}, \quad \forall n \geq n_1;$$

(2) $\exists n_m > n_{m-1}$ 使得

$$P\left(|\xi_n - \xi| > \frac{1}{m}\right) < \frac{1}{2^m}, \quad \forall n \geq n_m.$$

(3) $\forall \varepsilon, \exists m$ 使得 $\frac{1}{m} < \varepsilon$. 当 $k \geq m$ 时, $\frac{1}{k} < \varepsilon$, $n_k \geq n_m$,

$$P(|\xi_{n_k} - \xi| > \varepsilon) \leq P\left(|\xi_{n_k} - \xi| > \frac{1}{k}\right) < \frac{1}{2^k}.$$

(4) $\sum_k P(A_{k,\varepsilon}) < \infty \Rightarrow P(A_{k,\varepsilon}, \text{ i.o.}) = 0$.

- $\xi_n \xrightarrow{P} \xi$ iff $\forall \{n_k\}, \exists \{n_{k_i}\}$ 使得 $\xi_{n_{k_i}} \xrightarrow{\text{a.s.}} \xi$.

⇒: 已证

- ⇐: 反证法: 否则,

$$\exists \varepsilon, \delta > 0, n_1 < n_2 < \dots \text{ 使得 } P(|\xi_{n_k} - \xi| > \varepsilon) > \delta.$$

- 对于 $\{\xi_{n_k}, k \geq 1\}$, 不存在收敛子列:

否则 $\xi_{n_{k_i}} \xrightarrow{\text{a.s.}} \xi$,

于是 $\xi_{n_{k_i}} \xrightarrow{P} \xi$. 与 $P(|\xi_{n_{k_i}} - \xi| > \varepsilon) > \delta$ 矛盾!

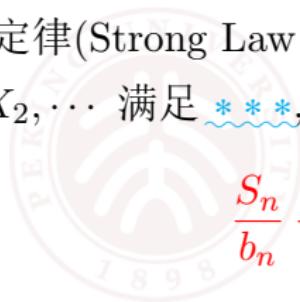
- 总结: $\xi_n \xrightarrow{\text{a.s.,r}} \eta \Rightarrow \xi_n \xrightarrow{P} \eta$, 反之不成立.

各种版本的SLLN.

强大数定律(Strong Law of Large Numbers, SLLN)形式:

若 X_1, X_2, \dots 满足 ***, 则

$$\frac{S_n}{b_n} - a_n \xrightarrow{\text{a.s.}} 0. \quad (5.4.28)$$



定理 ((Borel-)Cantelli's SLLN, 定理5.4.2, 习题五、46)

若 X_1, X_2, \dots 相互独立, $E(X_i - EX_i)^4 \leq M, \forall i$, 则

$$\frac{S_n - ES_n}{n} \xrightarrow{\text{a.s.}} 0.$$

- 不妨设 $EX_i = 0$. $A_n = \left\{ \left| \frac{S_n}{n} \right| \geq \varepsilon \right\}$. 往证 $\sum_n P(A_n) < \infty$.
- 将证 $ES_n^4 \leq 3n^2M, \forall n$. 于是,

$$P(A_n) = P(S_n^4 \geq (n\varepsilon)^4) \leq \frac{1}{n^4\varepsilon^4} ES_n^4 \leq \frac{3M}{\varepsilon^4} \cdot \frac{1}{n^2}.$$

(1) $ES_n^4 = \sum_{i,j,k,\ell=1}^n EX_i X_j X_k X_\ell,$
 $EX_r^4, EX_r^3 X_s, EX_r^2 X_s^2, EX_r^2 X_s X_t, EX_r X_s X_t X_u.$

(2) $E\star = 0, EX_r^4 \leq M,$
 $EX_r^2 X_s^2 = EX_r^2 \times EX_s^2 \leq \sqrt{EX_r^4} \sqrt{EX_s^4} \leq M.$

(3) 因此,

$$ES_n^4 \leq nM + C_n^2 C_4^2 M \leq 3n^2 M.$$

定理

若 $X = X_1, X_2, \dots$ 独立同分布, $EX^2 < \infty$, 则

$$\frac{1}{n} S_n \xrightarrow{\text{a.s.}} EX.$$

- 不妨设 $EX = 0, EX^2 = 1$. 令 $A_n = \left\{ \left| \frac{S_n}{n} \right| > \varepsilon \right\}$. 则

$$P(A_n) \leq \frac{1}{\varepsilon^2} E \left(\frac{S_n}{n} \right)^2 = \frac{1}{\varepsilon^2} \frac{n EX^2}{n^2} = \frac{1}{\varepsilon^2} \cdot \frac{1}{n}.$$

- $\frac{S_{m^2}}{m^2} \xrightarrow{\text{a.s.}} 0$:

$$\sum_m P(A_{m^2}) < \infty \Rightarrow P(A_{m^2} \text{ i.o.}) = 1.$$

- $|S_n| \leq |S_{m^2}| + T_m, \quad \forall m^2 \leq n < (m+1)^2 = m^2 + 2m + 1.$

$$T_m := \max_{1 \leq k \leq 2m} |S_k^{(m)}|,$$

其中, $S_k^{(m)} = X_{m^2+1} + \cdots + X_{m^2+k}$.

- $\frac{T_m}{m^2} \xrightarrow{\text{a.s.}} 0$. 于是 $\frac{S_n}{n} \xrightarrow{\text{a.s.}} 0$.

$$P\left(\left|\frac{S_{k \atop m^2}^{(m)}}{m^2}\right| > \varepsilon\right) \leq \frac{1}{\varepsilon^2} \frac{k}{m^4},$$

$$\Rightarrow P\left(\frac{|T_m|}{m^2} > \varepsilon\right) \leq 2m \cdot \frac{1}{\varepsilon^2} \frac{2m}{m^4} = \frac{4}{\varepsilon^2} \cdot \frac{1}{m^2}.$$

定理 (Kolmogorov's SLLN, 定理5.4.4)

假设 $X = X_1, X_2, \dots$ 独立同分布, $E|X| < \infty$, 则

$$\frac{S_n}{n} \xrightarrow{\text{a.s.}} EX.$$

- 时间平均= 空间平均, (期望的含义).
- 若 X_1, X_2, \dots 独立, 且 $\frac{S_n}{n} \xrightarrow{\text{a.s.}} Y$, 则 Y 退化.
- 若 i.i.d. 且 $\frac{S_n}{n} \xrightarrow{\text{a.s.}} a \in \mathbb{R}$, 则 $a = EX$.

(1) 在 $A = \{\lim_n \frac{S_n}{n} \exists\}$ 上,

$$\frac{S_{n-1}}{n-1} - \frac{S_n}{n} = \frac{S_{n-1}}{(n-1)n} - \frac{X_n}{n}, \quad \text{故 } \frac{X_n}{n} \rightarrow 0.$$

(2) 习题五、45. 假设 i.i.d. 则

$$\frac{X_n}{n} \xrightarrow{\text{a.s.}} 0 \quad \text{iff} \quad E|X| < \infty.$$

定理

假设 $X = X_1, X_2, \dots$ 独立同分布, $EX = \infty$, 则

$$\frac{S_n}{n} \xrightarrow{\text{a.s.}} \infty.$$

- $X = X^+ - X^-$. 不妨设 $X \geq 0$.

- 给定 $M > 0$, 令 $T_n = X_1 \wedge M + \cdots + X_n \wedge M$.

$$\frac{T_n}{n} \xrightarrow{\text{a.s.}} E(X \wedge M).$$

- $S_n \geq T_n$: $\forall M > 0$,

$$\liminf_{n \rightarrow \infty} \frac{S_n}{n} \geq E(X \wedge M), \quad \text{a.s.}$$

即, $\Omega_M = \{\star\star\}$, 则 $P(\Omega_M) = 1$.

- $P(\Omega_M) = 1$. 令 $\hat{\Omega} = \bigcap_{M=1}^{\infty} \Omega_M$. 则 $P(\hat{\Omega}) = 1$, 且在 $\hat{\Omega}$ 上,

$$\liminf_{n \rightarrow \infty} \frac{S_n}{n} \geq E(X \wedge M), \quad \forall M.$$

- $E(X \wedge M) \rightarrow \infty$:

$$\begin{aligned} E(X \wedge M) &= \int_0^{\infty} P(X \wedge M > x) dx \\ &= \int_0^M P(X > x) dx \rightarrow \infty. \end{aligned}$$

i.i.d. 序列总结:

- EX 有意义, 则

$$\frac{S_n}{n} \xrightarrow{\text{a.s.}} EX.$$

- $EX^\pm = \infty$: 柯西分布(例3.3.5),

$$\frac{S_n}{n} \stackrel{d}{=} X_1.$$

SLLN的应用.

例5.3.1. 样本均值& 样本方差.

- 数据: X_1, X_2, \dots 独立同分布.
- 样本均值:

$$\bar{X} = \frac{1}{n} S_n \xrightarrow{\text{a.s.}} EX.$$

- 样本方差: $\frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2 \approx$

$$\begin{aligned} \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X})^2 &= \frac{1}{n} \sum_{i=1}^n \textcolor{red}{X_i^2} - \textcolor{blue}{\bar{X}^2} \\ &\xrightarrow{\text{a.s.}} \textcolor{red}{EX^2} - (\textcolor{blue}{EX})^2 = \text{var}(X). \end{aligned}$$

例5.3.2. $\int_0^1 f(x)dx = E \mathbf{f}(\mathbf{U}).$

- SLLN:

$$\frac{1}{n} \sum_{i=1}^n \mathbf{f}(\mathbf{U}_i) \xrightarrow{\text{a.s.}} \int_0^1 f(x)dx.$$

- 高维:

$$\int_0^1 \cdots \int_0^1 f(\mathbf{x}_1, \dots, \mathbf{x}_n) d\vec{x} = E f(\mathbf{U}_1, \dots, \mathbf{U}_n).$$

习题五、50. 假设 $f, g : [0, 1] \rightarrow \mathbb{R}$, $0 \leq f < Cg$. 求

$$\lim_{n \rightarrow \infty} \int_0^1 \cdots \int_0^1 \frac{f(x_1) + \cdots + f(x_n)}{g(x_1) + \cdots + g(x_n)} dx_1 \cdots dx_n.$$

- 记 $X_i = f(U_i)$, $Y_i = g(U_i)$; 部分和: S_n , T_n . 令

$$W_n = \frac{S_n}{T_n}. \quad \text{再次强调: } EW_n \neq \frac{ES_n}{ET_n}.$$

- SLLN:

$$W_n = \frac{S_n/n}{T_n/n} \xrightarrow{\text{a.s.}} \frac{EX}{EY} = \frac{\int_0^1 f(x)dx}{\int_0^1 g(x)dx} = w.$$

- 有界收敛定理:

$$EW_n \rightarrow Ew = w.$$

§5.2 之依分布收敛, §5.3 之中心极限定理

- 权重的收敛.
- 例5.2.1: $\xi_n = \frac{1}{n}$, $\xi = 0$.

$$F_{\xi_n}(0) = 0, \quad F_\xi(0) = 1.$$



- 如果 $\forall x \in C(F_\xi)$ 都有

$$\lim_{n \rightarrow \infty} F_{\xi_n}(x) = F_\xi(x).$$

那么, 称 ξ_n 依分布收敛于 ξ , 记为 $\xi_n \xrightarrow{d} \xi$. (定义5.2.2 & 5.2.1)

- $\mathbb{R} \setminus C(F_\xi)$ 可数.

依分布收敛与依概率收敛.

- $\xi_n \xrightarrow{P} \xi \Rightarrow \xi_n \xrightarrow{d} \xi$. (定理5.2.6)

(1) $P(X \leq x) = P(X \leq x, Y \leq x) + P(\textcolor{red}{X} \leq x < \textcolor{blue}{Y})$:

$$|P(X \leq x) - P(Y \leq x)| \leq P(\textcolor{red}{X} \leq x < \textcolor{blue}{Y}) + P(\textcolor{red}{Y} \leq x < \textcolor{blue}{X}).$$

(2) 若 $|X - Y| < \varepsilon$ 且 $\textcolor{red}{X} \leq x < \textcolor{blue}{Y}$, 则 $x < \textcolor{blue}{Y} < x + \varepsilon$:

$$|P(X \leq x) - P(Y \leq x)| \leq P(|X - Y| \geq \varepsilon) + P(|Y - x| < \varepsilon).$$

(3) $|F_{\xi_n}(x) - F_\xi(x)| \leq P(|\xi_n - \xi| \geq \varepsilon) + P(|\xi - x| < \varepsilon)$.

- “ $\xi_n \xrightarrow{d} \xi \Rightarrow \xi_n \xrightarrow{P} \xi$ ” 不成立. (例5.2.3)
- $\xi_n \xrightarrow{d} C \Rightarrow \xi_n \xrightarrow{P} C$. (定理5.2.7)

例, $C = 0$:

$$P(|\xi_n| \leq \varepsilon) \geq F_{\xi_n}(\varepsilon) - F_{\xi_n}(-\varepsilon) \rightarrow F_\xi(\varepsilon) - F_\xi(-\varepsilon) = 1.$$

- $\xi_n \xrightarrow{d} \xi \Rightarrow \hat{\xi}_n \xrightarrow{\text{a.s.}} \hat{\xi}$, 其中

$$\hat{\xi}_n := F_{\xi_n}^{-1}(U) \stackrel{d}{=} \xi_n, \quad \hat{\xi} := F_\xi^{-1}(U) \stackrel{d}{=} \xi.$$

(证明不要求)

- 有界收敛定理: ξ_n 有界, $\xrightarrow{d} \xi$, 则 $E\xi_n \rightarrow E\xi$.

依分布收敛的等价条件.

$\xi_n \xrightarrow{d} \xi$ iff $Ef(\xi_n) \rightarrow Ef(\xi)$, $\forall f \in \mathbb{F}$.

- \mathbb{F}_1 : $f = 1_{(-\infty, b]}$, 其中 $b \in C(F_\xi)$. (定义)
- \mathbb{F}'_1 : $f = 1_{(-\infty, b]}$, 其中 $b \in C(F_\xi)$ 的某稠子集.
- \mathbb{F}_2 : $f = 1_{(a, b]}$, 其中, $a, b \in C(F_\xi)$.
- \mathbb{F}'_2 : 阶梯函数

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$$f = \sum_{i=1}^n \lambda_i \cdot 1_{(a_i, b_i]},$$

其中, $a_i, b_i \in C(F_\xi)$, $1 \leq i \leq n$.

- \mathbb{F}_3 : $f : \mathbb{R} \rightarrow \mathbb{R}$ 有界 连续. \Rightarrow : (假设 \xrightarrow{d} , 定理5.2.2 & 5.2.3).

(1) 取 a 使得 $\pm a \in C(F_\xi)$,

$$P(|\xi| \geq a) \leq \varepsilon \Rightarrow P(|\xi_n| \geq a) \leq 2\varepsilon, \quad \forall n \geq N.$$

(2) 取阶梯函数 g :

$$|f(x) - g(x)| \leq \varepsilon, \quad \forall |x| \leq a; \quad g(x) = 0, \quad |x| > a.$$

$$(3) \quad Ef(\xi_n) \cdot 1_{\{|\xi_n| \leq a\}} - Ef(\xi) \cdot 1_{\{|\xi| \leq a\}} = r_n + s_n + t_n:$$

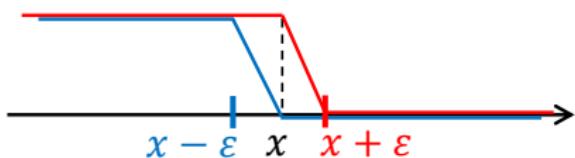
$$r_n = \star - Eg(\xi_n) \cdot 1_{\{|\xi_n| \leq a\}}, \quad t_n = Eg(\xi) \cdot 1_{\{|\xi| \leq a\}} - \star,$$

$$s_n = Eg(\xi_n) \cdot 1_{\{|\xi_n| \leq a\}} - Eg(\xi) \cdot 1_{\{|\xi| \leq a\}} = Eg(\xi_n) - Eg(\xi).$$

$$(4) \quad |Ef(\xi_n) - Ef(\xi)| \leq 3M\varepsilon + \varepsilon + |s_n|.$$

- \mathbb{F}_3 : $f : \mathbb{R} \rightarrow \mathbb{R}$ 有界连续. \Leftarrow :

(1) $g \leqslant 1_{(-\infty, x]} \leqslant f$.



(2) $1_{(-\infty, x]} \leqslant f$:

$$F_{\xi_n}(x) \leqslant Ef(\xi_n) \rightarrow Ef(\xi) \leqslant F_\xi(x + \varepsilon).$$

(3) $g \leqslant 1_{(-\infty, x]}$:

$$F_{\xi_n}(x) \geqslant Eg(\xi_n) \rightarrow Eg(\xi) \geqslant F_\xi(x - \varepsilon).$$

(4) $x \in C(F_\xi)$:

$$\lim_{\varepsilon \searrow 0} F_\xi(x + \varepsilon) = \lim_{\varepsilon \searrow 0} F_\xi(x - \varepsilon) = F_\xi(x).$$

- \mathbb{F}_4 : $f = \text{三角函数}$

$$\cos(tx), \quad \sin(tx), \quad t \in \mathbb{R}.$$

- 定理5.2.4 & 5.2.5 (证明不要求).

$\xi_n \xrightarrow{d} \xi$ iff 特征函数收敛: $f_{\xi_n}(t) \rightarrow f_\xi(t), \forall t$.

- 定理5.2.5, §4.5 连续性定理: 若

$$f_{\xi_n}(t) \rightarrow f(t), \quad \forall t, \quad \text{且 } f \text{ 在 } t = 0 \text{ 连续.}$$

则, f 是特征函数, ξ_n 依分布收敛. (证明不要求)

极限定理.

习题五、35. 当 $\delta_n := \max_{1 \leq k \leq n} p_k^{(n)} \rightarrow 0$, $\lambda_n := \sum_{k=1}^n p_k^{(n)} \rightarrow \lambda$ 时,

$$B(n, \vec{p})^{(n)} := B(1, p_1^{(n)}) * \cdots * B(1, p_n^{(n)}) \xrightarrow{d} P(\lambda).$$

- $B(n, \vec{p}^{(n)})$ 的特征函数:

$$\begin{aligned} f_n(t) &= \prod_{k=1}^n \left((1 - p_k^{(n)}) + p_k^{(n)} e^{it} \right) \\ &= \prod_{k=1}^n (1 + p_k^{(n)} z), \quad \text{其中 } z = e^{it} - 1. \end{aligned}$$

- $P(\lambda_n) \rightarrow P(\lambda)$: 令 $g_n(t)$ 表示 $P(\lambda_n)$ 的特征函数, 则

$$g_n(t) = \prod_{k=1}^n e^{p_k^{(n)} z} = e^{\lambda_n z} \rightarrow e^{\lambda z}.$$

- $a_k = 1 + p_k^{(n)} z$, $b_k = e^{p_k^{(n)} z}$,

$$\|\textcolor{red}{a_k}\| = \|1 + p_k^{(n)} z\| = \|(1 - p_k^{(n)}) \cdot 1 + p_k^{(n)} \cdot e^{it}\| \leq 1.$$

$$\|\textcolor{blue}{b_k}\| = \|e^{p_k^{(n)} z}\| = \|e^{p_k^{(n)}(\cos t - 1) + i p_k^{(n)} \sin t}\| = e^{p_k^{(n)}(\cos t - 1)} \leq 1.$$

- $f_n(t) = \prod_{k=1}^n \textcolor{red}{a_k}$, $g_n(t) = \prod_{k=1}^n \textcolor{blue}{b_k}$, 故

$$\|f_n(t) - g_n(t)\| \leq \sum_{k=1}^n \|a_k - b_k\|. \quad (5.5.13)$$

- 归纳法:

$$\left\| \prod_{k=1}^n a_k - \prod_{k=1}^{n-1} a_k \times \textcolor{blue}{b_n} \right\| \leq 1 \times \|a_n - b_n\|;$$

$$\left\| \prod_{k=1}^{n-1} a_k \times b_n - \prod_{k=1}^{n-1} b_k \times \textcolor{blue}{b_n} \right\| \leq \left\| \prod_{k=1}^{n-1} a_k - \prod_{k=1}^{n-1} b_k \right\| \times 1.$$

- 记 $s = p_k^{(n)} z$, 则 $a_k = 1 + s$, $b_k = e^s$, 故 $\|a_k - b_k\| \leq \|s\|^2$:

$$\begin{aligned}\|a_k - b_k\| &= \left\| \sum_{\ell=2}^{\infty} \frac{s^\ell}{\ell!} \right\| \leq \sum_{\ell=2}^{\infty} \frac{\|s\|^\ell}{\ell!} \leq \sum_{\ell=2}^{\infty} \frac{\|s\|^2}{\ell!} \\ &\leq \|s\|^2 (e - 2) \leq (p_k^{(n)})^2 \|z\|^2.\end{aligned}$$

- 注: 以上的估计并不精细.
- 故,

$$\|f_n(t) - g_n(t)\| \leq \sum_{k=1}^n \delta_n p_k^{(n)} \|z\|^2 \leq 4\delta_n \lambda_n \rightarrow 0.$$

定理5.3.1. 设 $X = X_1, X_2, \dots$ i.i.d., $E|X| < \infty$, 则 $\frac{S_n}{n} \xrightarrow{P} EX$.

证明 方法II

- 不妨设 $EX = 0$. 只需证

$$f_n(t) = f_{\frac{S_n}{n}}(t) \rightarrow 1 = e^{it \times 0}.$$

- 记 $f = f_X$. 于是

$$f_n(t) = E \exp \left\{ it \frac{X_1 + \dots + X_n}{n} \right\} = \left(E e^{i \frac{t}{n} X} \right)^n = f \left(\frac{t}{n} \right)^n.$$

- 注: 根据有界收敛定理, $E e^{i \frac{t}{n} X} \rightarrow e^{i \times 0} = 1$.
往估计收敛速度.

- $e^{is} = 1 + is + \delta(s)$, $\|\delta(s)\| \leq \min\left\{2|s|, \frac{1}{2}s^2\right\}$. (引理5.5.1)
- $f(\frac{t}{n})$:

$$E\left(1 + i\frac{t}{n}X + \delta\left(\frac{t}{n}X\right)\right) = 1 + 0 + E\delta\left(\frac{t}{n}X\right).$$

- $\|E\delta(\frac{t}{n}X)\| = o(\frac{1}{n})$:

$$\begin{aligned} \left\|E\delta\left(\frac{t}{n}X\right)\right\| &= E\left\|\delta\left(\frac{t}{n}X\right)\right\| \cdot 1_{\{|X|>M\}} + E\left\|\delta\left(\frac{t}{n}X\right)\right\| \cdot 1_{\{|X|\leq M\}} \\ &\leq \frac{2t}{n}E|X|1_{\{|X|>M\}} + \frac{t^2}{2n^2}M^2. \end{aligned}$$

- $f(\frac{t}{n})^n = (1 + o(\frac{1}{n}))^n \rightarrow 1$.
- $\frac{1}{n}S_n \xrightarrow{d,P} 0$.

定理 (Lindeberg-Levy 中心极限定理, 定理5.4.4)

假设 X_1, X_2, \dots 独立同分布. $0 < \text{var}(X) < \infty$. 则

$$S_n^* \xrightarrow{d} Z \sim N(0, 1).$$

- 不妨设 $EX = 0, EX^2 = 1$. 记 $f = f_{X_1}$, 则

$$f_{S_n^*}(t) = E \exp \left\{ i \textcolor{blue}{t} \frac{S_n}{\sqrt{n}} \right\} = f \left(\frac{\textcolor{blue}{t}}{\sqrt{n}} \right)^n.$$

- (记 $s = \frac{t}{\sqrt{n}}$.) 往证

$$f(s) = E e^{isX} = 1 - \frac{1}{2}s^2 + o(s^2).$$

$$\text{于是 } f_{S_n^*}(t) = \left(1 - \frac{t^2}{2n} + o\left(\frac{1}{n}\right) \right)^n \rightarrow e^{-t^2/2}.$$

引理5.5.1. (记 $y = sX$.)

$$|e^{iy} - 1| \leq |y|, \quad |e^{iy} - 1 - iy| \leq \frac{|y|^2}{2} \quad (5.5.10),$$

$$|e^{iy} - 1 - iy + \frac{y^2}{2}| \leq \frac{|y|^3}{6} \quad (5.5.11).$$

- $e^{iy} = 1 + iy + \frac{1}{2}(iy)^2 + \varphi(y).$

$$\begin{aligned}\varphi(y) &= e^{iy} - 1 - iy - \frac{1}{2}(iy)^2 \\&= \int_0^y ie^{iz} dz - \int_0^y i \cdot 1 dz - \frac{1}{2}(iy)^2 \\&= \int_0^y i \underbrace{\int_0^z ie^{iw} dw dz}_{\text{交换积分次序}} - \int_0^y \int_0^z i^2 dw dz \\&= \int_0^y \int_0^z i^2(e^{iw} - 1) dw dz = \int_0^y \int_0^z \int_0^w i^3 e^{iu} du dw dz.\end{aligned}$$

- $\|\varphi(y)\| \leq y^2;$ $\|\varphi(y)\| \leq \frac{|y|^3}{6}:$

$$\varphi(y) = \int_0^y \int_0^z i^2(e^{iw} - 1) dw dz = \int_0^y \int_0^z \int_0^w i^3 e^{iu} du dw dz.$$

- 取 $y = sX$. 则

$$e^{isX} = 1 + isX + \frac{1}{2}(isX)^2 + \varphi(sX),$$

$$f(s) = Ee^{isX} = 1 + 0 - \frac{1}{2}s^2 + E\varphi(sX).$$

- 由 $\|\varphi(y)\| \leq y^2$, $\|\varphi(y)\| \leq \frac{|y|^3}{6}$ 可推出 $\|E\varphi(sX)\| = o(s^2)$:

$$\begin{aligned}\|E\varphi(sX)\| &\leq E\|\varphi(sX)\| \cdot 1_{\{|X|>M\}} + E\|\varphi(sX)\| \cdot 1_{\{|X|\leq M\}} \\ &\leq E(sX)^2 \cdot 1_{\{|X|>M\}} + E\frac{|sX|^3}{6} \cdot 1_{\{|X|\leq M\}} \\ &\leq s^2 EX^2 \cdot 1_{\{|X|>M\}} + |s|^3 \cdot \frac{M^3}{6}.\end{aligned}$$

- 注:

$$f(s) = 1 - \frac{1}{2}s^2 + E\varphi(sX) = 1 - \frac{1}{2}s^2 + o(s^2).$$

- 固定 t . 取 $s = \frac{t}{\sqrt{n}}$, 令 $n \rightarrow \infty$. 则

$$f\left(\frac{t}{\sqrt{n}}\right) = 1 - \frac{t^2}{2n} + \varepsilon_n,$$

其中, $\varepsilon_n := E\varphi\left(\frac{t}{\sqrt{n}}X\right) = o\left(\frac{t^2}{\sqrt{n}^2}\right) = o\left(\frac{1}{n}\right)$.

- $S_n^* \xrightarrow{d} Z \sim N(0, 1)$:

$$f_{S_n^*}(t) = f\left(\frac{t}{\sqrt{n}}\right)^n = \left(1 - \frac{t^2}{2n} + \varepsilon_n\right)^n \rightarrow e^{-\frac{t^2}{2}}.$$

- 中心极限定理(Central Limit Theorem, CLT)形式:

若 X_1, X_2, \dots 满足 ***, 则

$$S_n^* \xrightarrow{d} Z \sim N(0, 1). \quad (5.1.11)$$

- Lindeberg-Levy 版本: *** = “i.i.d., $0 < \text{var}(X) < \infty$ ”.

- $\frac{1}{\sqrt{n}}(S_n - ES_n) \xrightarrow{d} \sigma Z \sim N(0, \sigma^2).$

- 若 $\vec{X} = \vec{X}_1, \vec{X}_2, \dots$ i.i.d., $\vec{\mu} = E\vec{X}$, $\Sigma = (\text{cov}(X_i, X_j))$. 则

$$\frac{1}{\sqrt{n}}(\vec{S}_n - n\vec{\mu}) \xrightarrow{d} \sqrt{\Sigma}\vec{Z} \sim N(\vec{0}, \Sigma).$$

- Berry-Esseen's bound: 假设 $E|X|^3 < \infty$. 那么,

$$|F_{S_n^*}(x) - \Phi(x)| \leq \frac{3E|X^*|^3}{\sqrt{n}}, \quad \forall x.$$

(了解即可.)

CLT的应用.

例: 经验分布函数.

$$\text{SLLN: } F_n(x) = \frac{1}{n} \sum_{i=1}^n 1_{\{X_i \leq x\}} \xrightarrow{\text{a.s.}} F(x).$$

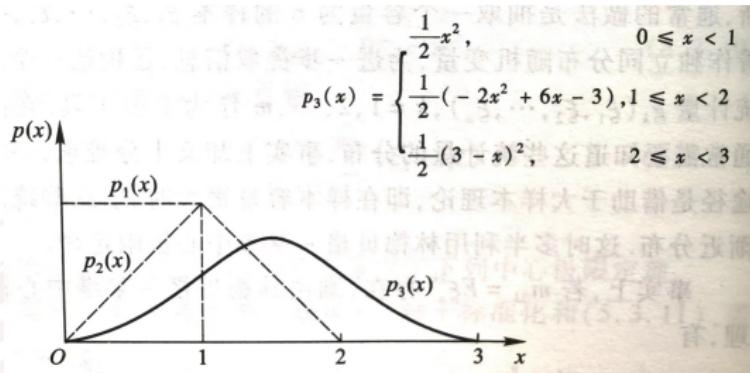
- CLT:

$$\begin{aligned}\sqrt{n}(F_n(x) - F(x)) &\xrightarrow{d} \sqrt{F(x)(1-F(x))} Z \\ &\sim N(0, F(x)(1-F(x))).\end{aligned}$$

- (了解即可) $D_n = \max_x |F_n(x) - F(x)|$:

$$\sqrt{n}D_n \xrightarrow{d} \xi, \quad F_\xi(x) = \sum_{k=-\infty}^{\infty} (-1)^k e^{-2k^2x^2}, \quad \forall x > 0.$$

- 例5.3.4. X_1, X_2, \dots i.i.d., $\sim U(0, 1)$.



- 将 S_n^* “视为” $\sim N(0, 1)$:

$$P(S_n \leq x) = P(S_n^* \leq x^*) \approx \Phi(x^*) = p, \quad x^* = \frac{x - E S_n}{\sqrt{\text{var}(S_n)}}.$$

- n, x, p 满足关系式:

$$P(S_n \leq x) = p.$$

例5.3.5. 在近似计算中保留5位小数. 考虑 10^4 个数, 近似值之和与真值之和的误差不超过 10^{-3} 的概率.

- 近似值与真值的误差 $X \sim U(-0.5 \times 10^{-5}, 0.5 \times 10^{-5})$.
- $n = 10^4$, $x = 10^{-3}$, 求 p .
- 目标:

$$P(|S_n| \leq 10^{-3}) = P(|S_n^*| \leq x^*) = p,$$

其中, $x^* = \frac{10^{-3}}{\sqrt{n \cdot \text{var}(X)}} = \frac{10^{-3}}{\sqrt{10^4 \cdot 10^{-5} \sqrt{\frac{1}{12}}}} = 2\sqrt{3} = 3.46 \dots$.

- 查表: $\Phi(3.46) = 99.97299\%$,
- 故 $p \geq 2\Phi(3.46) - 1 = 99.94598\%$.

例2.4.4. 人寿保险. 死亡率 $p = 0.005$, 参保数 $n = 10^4$. 估算:

$$P(\text{死亡人数} = 40), P(\text{死亡人数} \leq 70).$$

- $P(A) = C_{10000}^{40} p^{40} q^{9960} = \underline{0.0214}$.

- $x^* = \frac{x-np}{\sqrt{npq}}$:

$$P(A) \approx \varphi(40^*) \frac{1}{\sqrt{npq}} = \underline{0.0207}.$$

- 连续修正: $P(A) = P(39.5 \leq S_n \leq 40.5)$,

$$x_1 = 39.5^* \approx -1.49, x_2 = 40.5^* \approx -1.35:$$

$$P(A) = P(39.5 \leq S_n \leq 40.5) = P(x_1 \leq S_n^* \leq x_2)$$

$$\approx 0.93189 - 0.91149 = \underline{0.0204}.$$

- 连续修正: $P(B) = P((-0.5)^* \leq S_n^* \leq (70.5)^*) \approx 0.998193$.

- 如果不连续修正: $P(0^* \leq S_n \leq 70^*) \approx 0.997744$.

例2.4.6. 车间有200台车床, 每台的功率为1 千瓦, 开动频率为60%. 要求正常生产的把握至少为99.9%, 问: 需多少千瓦电力?

- 已知 $n = 200$, $p = 99.9\%$, 求 x .
- $P(X = 1) = 1 - P(X = 0) = 0.6$. $n = 200$. $p = 99.9\%$.
- 目标: $x^* = \frac{x - 200 * 0.6}{\sqrt{200 * 0.6 * 0.4}}$,

$$P(S_n \leq x) = P(S_n^* \leq x^*) \geq 99.9\% = p.$$

- 查表: $\Phi(3.09) = 99.8999\%$, $\Phi(3.1) = 99.90324\%$.
- $\star \geq 3.1$, $x \geq 120 + 3.1\sqrt{48} \in (141, 142)$. 需142 千瓦.

例5.1.2. 往常的市场占有率为15%. 为保证调查结果与真实值的误差不超过0.01 的概率至少为95%. 至少需调查多少人?

- 已知 $x = 0.01$, $p = 0.95$, 求 n .
- $P(X = 1) = 1 - P(X = 0) = q = 0.15$. 目标:

$$P\left(\left|\frac{S_n}{n} - q\right| \leqslant 0.01\right) \geqslant 0.95.$$

$$P\left(|S_n^*| \leqslant \frac{0.01\sqrt{n}}{\sqrt{q(1-q)}} = x^*\right) \geqslant 0.95.$$

- $\Phi(x^*) \geqslant \frac{1}{2}(1 + 0.95) = 0.975$, 查表得 $x^* \geqslant 1.96$.
- $n \geqslant 196^2 q(1 - q)$, 至少调查4899 人.
- 若 q 未知, 则 $n \geqslant 196^2 \max_q q(1 - q) = 196^2 \cdot \frac{1}{4} = 9604$.

§5.5 中心极限定理

- 研究对象:

$$S_n^* = \sum_{k=1}^n \frac{X_k - \mu}{\sqrt{\text{var}(S_n)}} = \sum_{k=1}^n Y_k \stackrel{\text{iid}}{=} \text{时} \sum_{k=1}^n \frac{1}{\sqrt{n}} X_k^*.$$

- 假设 X_1, X_2, \dots 相互独立, $EX_k = \mu_k$, $\text{var}(X_k) = \sigma_k^2$. 记

$$B_n^2 = \sum_{k=1}^n \sigma_k^2.$$

- 目标:

$$S_n^* = \frac{S_n - \sum_{k=1}^n \mu_k}{B_n} = \sum_{k=1}^n \frac{X_k - \mu_k}{B_n} \xrightarrow{d} Z \sim N(0, 1).$$

$$S_n^* = \sum_{k=1}^n Y_k, \quad \text{其中, } Y_k := \frac{X_k - \mu_k}{B_n}.$$

- Lindeberg 条件 (5.5.2):

$$\sum_{k=1}^n E|Y_k|^2 \cdot 1_{\{|Y_k| > \varepsilon\}} \rightarrow 0, \quad \forall \varepsilon > 0.$$

- Lindeberg 条件 $\implies P(\max_{1 \leq k \leq n} |Y_k| > \varepsilon) \rightarrow 0 \quad (n \rightarrow \infty).$
- Feller 条件 (5.5.3): $\frac{\max_{1 \leq k \leq n} \sigma_k}{B_n} \rightarrow 0.$

$$\frac{\max_{1 \leq k \leq n} \sigma_k}{B_n} \rightarrow 0 \quad \text{iff} \quad B_n \rightarrow \infty \text{ 且 } \frac{\sigma_n}{B_n} \rightarrow 0. \quad (\text{定理5.5.1})$$

- Lindeberg-Feller CLT (定理5.5.2):

$$(5.5.2) \quad \text{iff} \quad S_n^* \xrightarrow{d} Z \sim N(0, 1) \quad \text{且 (5.5.3).}$$